

# TSDT14 Signal Theory

## Lecture 4 Estimation

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# DFT – Signal Analysis

Time-discrete signal with limited duration:

$$x[n] = 0 \text{ for } n \notin \{0, 1, \dots, N-1\}$$

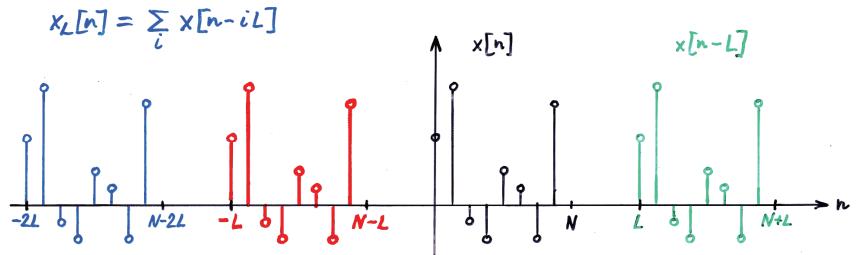
Fourier transform:  $X[\theta] = \sum_n x[n] e^{-j2\pi\theta n} = \sum_{n=0}^{N-1} x[n] e^{-j2\pi\theta n}$   
cont. w. period 1.

DFT of length  $L$ :  $X_L[k] = X[k/L] = \sum_{n=0}^{N-1} x[n] e^{-j2\pi\frac{k}{L}n} \text{ for } k \in \{0, 1, \dots, L-1\}$

IDFT (inverse):  $x_L[n] = \frac{1}{L} \cdot \sum_{k=0}^{L-1} X_L[k] e^{j2\pi\frac{k}{L}n} \Rightarrow x_L[n+L] = x_L[n]$   
since  $e^{j2\pi L n} = e^{j2\pi n} = 1$

Relation to  $x[n]$ :  $x_L[n] = \frac{1}{L} \sum_{k=0}^{L-1} \sum_{m=0}^{N-1} x[m] e^{-j2\pi\frac{k}{L}m} \cdot e^{j2\pi\frac{k}{L}n}$   
 $= \sum_{m=0}^{N-1} x[m] \cdot \underbrace{\frac{1}{L} \sum_{k=0}^{L-1} e^{-j2\pi(m-n)\frac{k}{L}}}_{\begin{cases} L, m-n=0 \bmod L \\ 0, \text{ elsewhere} \end{cases}} = \sum_{i=-\infty}^{\infty} x[n-i \cdot L]$

# DFT – Avoiding Aliasing



If  $L < N$ , then we get overlap and aliasing in the time domain.

Therefore: Demand  $L \geq N$ .

Note:  $X_L[k] = \sum_{n=0}^{N-1} x[n] e^{-j2\pi\frac{k}{L}n} = \sum_{n=0}^{L-1} x_L[n] e^{-j2\pi\frac{k}{L}n}$

# DFT – Periodic Convolution

We are used to:  $y[n] = (x * h)[n] \Leftrightarrow Y[\theta] = X[\theta] \cdot H[\theta]$

But we have:  $y[n] = x[n] \cdot h[n] \Leftrightarrow Y[\theta] = \int_0^1 X[\phi] H[\theta-\phi] d\phi$

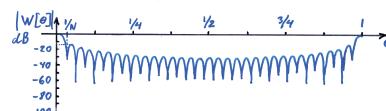
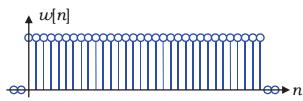
With DFT:  $\underline{Y_L[k]} = \underline{X_L[k]} \cdot \underline{H_L[k]} \Leftrightarrow$

$$\begin{aligned} y_L[n] &= IDFT \left\{ \underline{X_L[k]} \underline{H_L[k]} \right\} = \frac{1}{L} \sum_{k=0}^{L-1} \underline{X_L[k]} \underline{H_L[k]} e^{j2\pi\frac{k}{L}n} \\ &= \frac{1}{L} \sum_{k=0}^{L-1} \underline{X_L[k]} \cdot \sum_{m=0}^{L-1} \underline{h_L[m]} \cdot e^{-j2\pi\frac{k}{L}m} \cdot e^{j2\pi\frac{k}{L}n} \\ &= \sum_{m=0}^{L-1} \underline{h_L[m]} \cdot \underbrace{\frac{1}{L} \sum_{k=0}^{L-1} \underline{X_L[k]} e^{j2\pi\frac{k}{L}(n-m)}}_{x_L[n-m]} \\ &= \sum_{m=0}^{L-1} \underline{h_L[m]} x_L[n-m] \end{aligned}$$

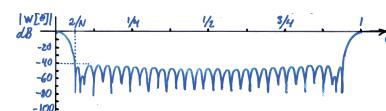
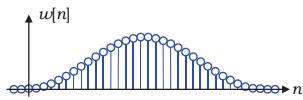
And also:  $y_L[n] = x_L[n] \cdot h_L[n] \Leftrightarrow \underline{Y_L[k]} = \frac{1}{L} \sum_{m=0}^{L-1} \underline{X_L[m]} \underline{H_L[k-m]}$

## Examples of Windows, $N=32$

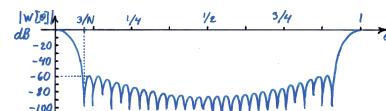
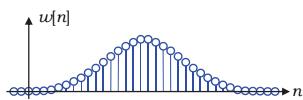
Rectangular window:  $w[n] = 1, n \in \{0, 1, \dots, N-1\}$



Hamming window:  $w[n] = 0.54 - 0.46\cos(\frac{2\pi N}{N-1}), n \in \{0, 1, \dots, N-1\}$

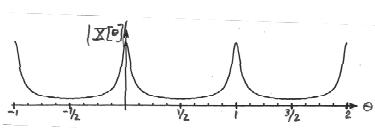
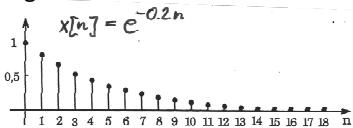


Blackman window:  $w[n] = 0.42 - 0.5\cos(\frac{2\pi N}{N-1}) + 0.08\cos(\frac{4\pi N}{N-1}), n \in \{0, 1, \dots, N-1\}$

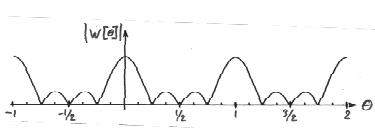
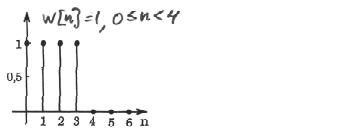


## Using Windows

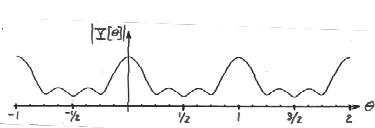
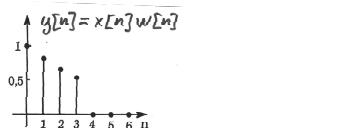
Signal:



Rectangular window:



Result:



## Ergodicity again

A WSS process:

$$m_{\bar{x}} = E\{\bar{x}(t)\}$$

Time average of one realization:

$$m_T = \frac{1}{2T} \int_{-T}^T x(t) dt$$

Time average of the process:

$$M_T = \frac{1}{2T} \int_{-T}^T \bar{x}(t) dt$$

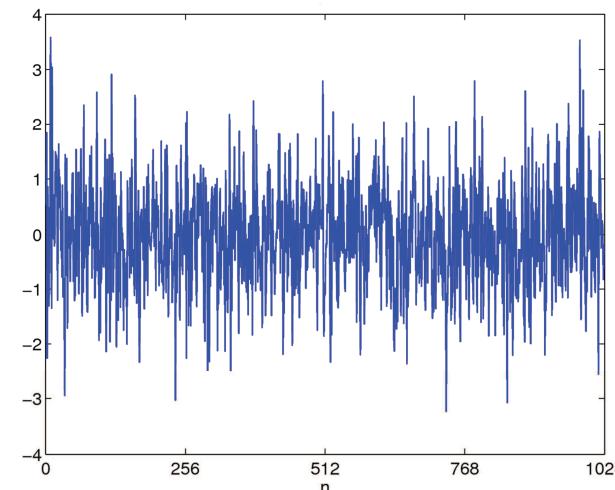
$$E\{M_T\} = E\left\{\frac{1}{2T} \int_{-T}^T \bar{x}(t) dt\right\} = \frac{1}{2T} \int_{-T}^T E\{\bar{x}(t)\} dt = \frac{1}{2T} \int_{-T}^T m_{\bar{x}} dt = m_{\bar{x}}$$

Definition: If  $\lim_{T \rightarrow \infty} E\{(M_T - m_{\bar{x}})^2\} = 0$  then  $\bar{x}(t)$  is said to be ergodic with respect to the mean, and we write

$$m_{\bar{x}} = \text{l.i.m. } \frac{1}{T} \int_{-T}^T \bar{x}(t) dt \quad (\text{limes in mean})$$

Interpretation: The time average of a realization is very close to the ensemble mean with a probability that is very close to 1 ( $\rightarrow 1, T \rightarrow \infty$ ).

## A Signal



## Estimating a General Parameter

$X[n]$ : An ergodic time-discrete process, with realization  $x[n]$ .

Estimation of a general parameter  $a_X$ , i.e. some ensemble average:

$$\hat{a}_X = g(x[0], x[1], \dots, x[N-1])$$

Corresponding stochastic variable:  $\hat{A}_X = g(X[0], X[1], \dots, X[N-1])$

Bias (difference from actual value):  $B = E\{\hat{A}_X\} - a_X$

Unbiased estimate:  $B = 0$

Asymptotically unbiased estimate:  $B \rightarrow 0$  when  $N \rightarrow \infty$

Variance:  $\sigma_{\hat{A}_X}^2 = E\{\hat{A}_X^2\} - E^2\{\hat{A}_X\}$

Quadratic error:  $\varepsilon^2 = \sigma_{\hat{A}_X}^2 + B^2$

Consistent estimate:  $\varepsilon^2 \rightarrow 0$  when  $N \rightarrow \infty$

## Blackman-Tukey's Estimate of the ACF 1(2)

ACF of a WSS process:  $r_X[k] = E\{X[n+k]X[n]\}$

Blackman-Tukey's method:  $\hat{r}_X[k] = \frac{1}{N-|k|} \sum_{n=0}^{N-|k|-1} x[n+|k|]x[n]$

Mean:  $E\{\hat{r}_X[k]\} = r_X[k]$

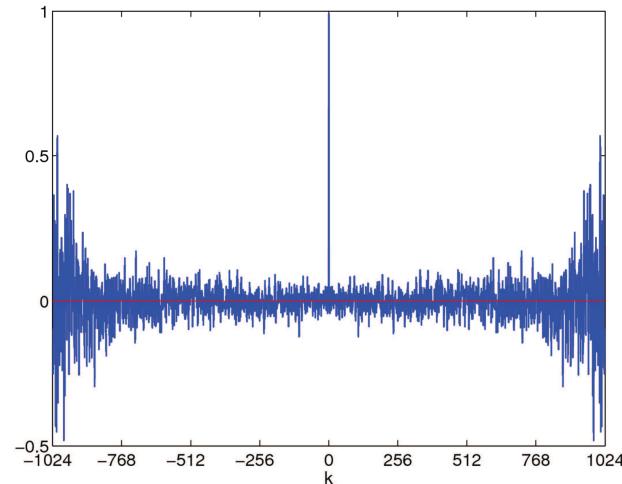
Variance, assuming a Gaussian process with mean zero:

$$|k| < N: \sigma_{\hat{r}_X}^2[k] = \frac{1}{N-|k|} \sum_{|m|<N-|k|} \left(1 - \frac{1}{N-|k|}\right) (r_X^2[m] + r_X[m+k]r_X[m-k])$$

Fixed  $k$ :  $\sigma_{\hat{r}_X}^2[k] \rightarrow 0$  when  $N \rightarrow \infty$

Problem: Large  $k$ , Ex.  $k = N-1$   $\sigma_{\hat{r}_X}^2[k] \rightarrow r_X^2[0]$  when  $N \rightarrow \infty$

## Blackman-Tukey's Estimate of the ACF 2(2)



## Bartlett's Estimate of the ACF 1(2)

Recall Blackman-Tukey:  $\hat{r}_X[k] = \frac{1}{N-|k|} \sum_{n=0}^{N-|k|-1} x[n+|k|]x[n]$

Bartlett's method:  $\hat{r}_X[k] = \frac{1}{N} \sum_{n=0}^{N-|k|-1} x[n+|k|]x[n]$

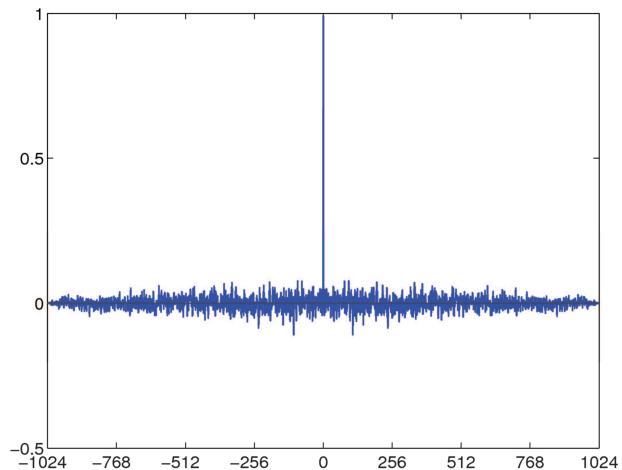
Mean:  $E\{\hat{r}_X[k]\} = \left(1 - \frac{|k|}{N}\right) r_X[k] \rightarrow r_X[k] \text{ when } N \rightarrow \infty$

Variance, assuming a Gaussian process with mean zero:

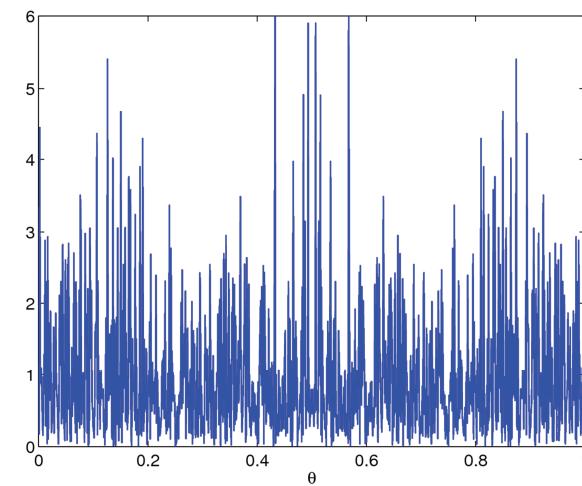
$$|k| < N: \sigma_{\hat{r}_X}^2[k] = \frac{1}{N} \sum_{|m|<N-|k|} \left(1 - \frac{|k|+|m|}{N}\right) (r_X^2[m] + r_X[m+k]r_X[m-k])$$

All  $k$ :  $\sigma_{\hat{r}_X}^2[k] \rightarrow 0$  when  $N \rightarrow \infty$

## Bartlett's Estimate of the ACF 2(2)



## Periodogram



## Estimating the PSD 1(2)

PSD of a WSS process:

$$R_X[\theta] = \mathcal{F}\{r_X[k]\}$$

Observed sequence:

$$x[n]$$

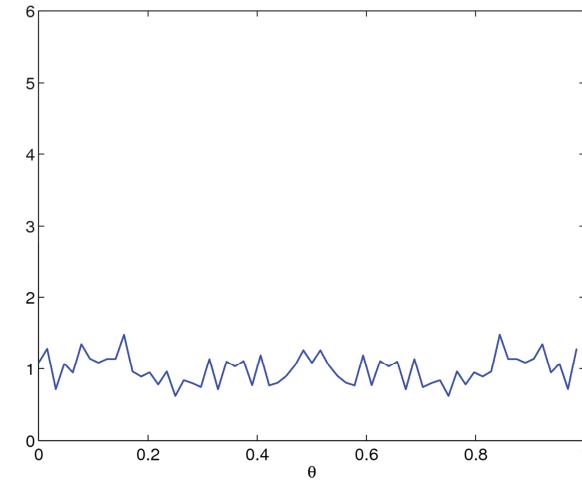
Considered part of sequence:

$$x_N[n] = \begin{cases} x[n], & 0 \leq n < N \\ 0 & \text{elsewhere} \end{cases}$$

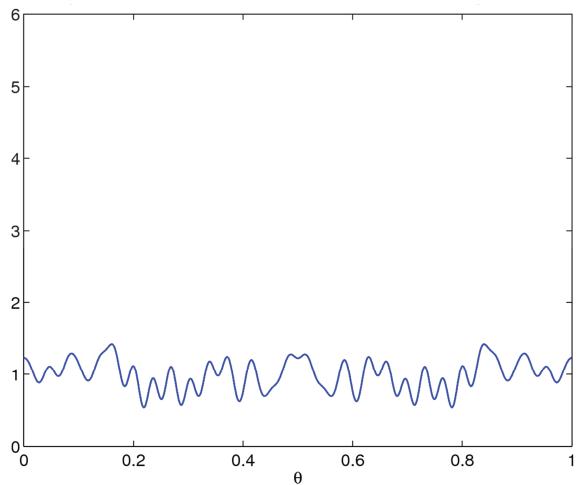
$$\text{Bartlett's estimate of the ACF: } \hat{r}_X[k] = \frac{1}{N} \sum_{m=-\infty}^{\infty} x_N[m+k]x_N[m]$$

$$\begin{aligned} \text{Estimate of the PSD: } & \hat{R}_X[\theta] = \mathcal{F}\{\hat{r}_X[k]\} \\ &= \sum_{k=-\infty}^{\infty} \frac{1}{N} \sum_{m=-\infty}^{\infty} x_N[m+k]x_N[m]e^{-j2\pi\theta k} = \frac{1}{N} \sum_{m=-\infty}^{\infty} x_N[m] \sum_{k=-\infty}^{\infty} x_N[m+k]e^{-j2\pi\theta k} \\ &= \frac{1}{N} X_N[\theta] \sum_{m=-\infty}^{\infty} x_N[m]e^{j2\pi\theta m} = \frac{1}{N} X_N[\theta] X_N^*[\theta] = \frac{1}{N} |X_N[\theta]|^2 \quad \text{Called periodogram} \end{aligned}$$

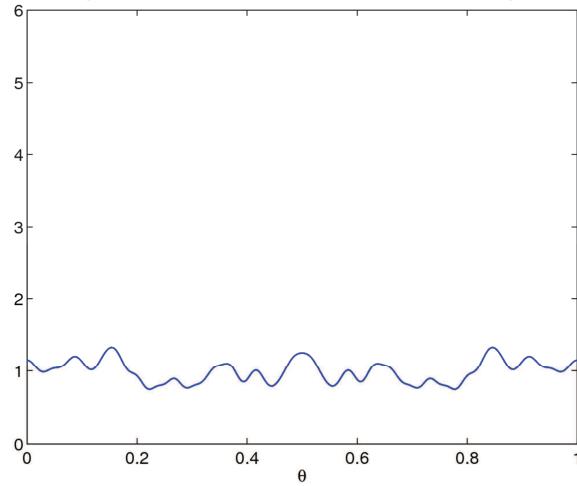
## Averaged Periodogram



## Estimated PSD using Rectangular Window

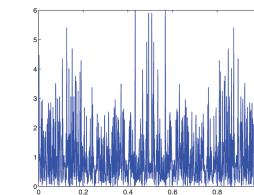


## Estimated PSD using Hamming Window

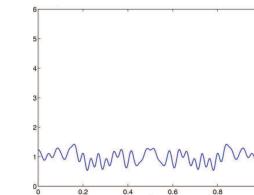


## Smoothing – Overview

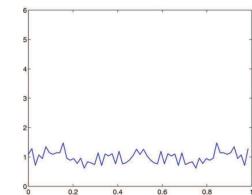
Raw periodogram



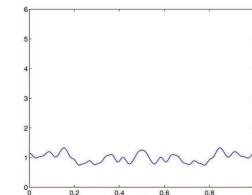
Smoothing - Rectangular



Averaged periodograms



Smoothing - Hamming



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